Next we will discuss two important tails:

$$\left\{ \begin{array}{ll} \theta_{0,n} \geq n(\mu + \varepsilon) & \mbox{ right tail} \\ \theta_{0.n} \leq n(\mu - \varepsilon) & \mbox{ left tail} \end{array} \right.$$

 $\theta = a, b, c$  or  $\phi$ , where  $a_{0,n}, b_{0,n}$  and  $c_{0,n}$  has been defined, while  $\phi_{0,n}$ , which is called "Face-face first-passage time", is defined as follows:

 $\phi_{0,n} := \inf\{T(\gamma): \gamma \text{ is a path from A to B}, \gamma \subseteq [0,n]^d\}.$ 

where A:= $\{0\} \times [0, n]^{d-1}$ , B:= $\{n\} \times [0, n]^{d-1}$ .

The examples of d = 2 and d = 3 can easily be imagined.

Theorem 1. minimum-cut=maximum flow.

Grimmett and Kesten (1984, PTRF) proved the following:

$$\lim_{n \to \infty} \frac{\phi_{0,n}}{n} = \mu \qquad a.s. \quad and \quad in \quad L_1$$

First we check the left tail.

**Theorem 2.(Kesten,1986)** Assume that for some r > 0,  $Ee^{rt(e)} < \infty$  and  $F(0) < p_c$ , then for all  $\varepsilon > 0$ , there exists constants  $c_1$  and  $c_2$ , such that

$$P(\theta_{0,n} \le n(\mu - \varepsilon)) \le c_1 \exp(-c_2 n).$$

 $\theta = a, b, c \text{ or } \phi.$ 

**Remark.** The assumption " $Ee^{rt(e)} < \infty$ " is not in Kesten(1986). In that paper it just assumed " $Et(e) < \infty$ ", but given the assumption " $Ee^{rt(e)} < \infty$ ", the proof will be much easier.

**Proof of Theorem 2 for**  $\theta = c$ 

$$P(c_{0,n} \le n(\mu - \varepsilon)) \le P(c_M(a_0) + c_M(a_1) + \dots + c_M(a_Q) \le n(\mu - \varepsilon)) \\ \le P(c'_M(a_0) + c'_M(a_1) + \dots + c'_M(a_{Q-1}) \le n(\mu - \varepsilon))$$
 (by Lemma 2)  
$$\le P(\sum_{i=0}^{Q-1} c'_M(a_i) \le QM(\mu - \varepsilon))$$

Since  $\lim_{M \longrightarrow \infty} \frac{c_{0,M}}{M} = \mu$ . Take M, such that  $Ec'_{0,M} \ge M(\mu - \frac{\varepsilon}{2})$ .

$$\sum_{i=0}^{Q-1} [c'_M(a_i) - Ec'_M(a_i)] \le QM(\mu - \varepsilon) - QM(\mu - \frac{\varepsilon}{2}) = -QM(\frac{\varepsilon}{2}).$$

Let  $X_i = c'_M(a_i) - Ec'_M(a_i)$ , then  $EX_i = 0$ ,  $\int_0^1 e^{\beta x} dF_{X_1}(x) < \infty$ .

For  $\beta = \beta(M, F, d)$ . By a standard large deviation result,  $\exists c_1, c_2$ , such that

$$P(c_{0,n} \le n(\mu - \varepsilon)) \le P(\sum_{i=1}^{Q} X_i \le -QM(\frac{\varepsilon}{2})) \le c_1 \exp(-c_2 n).$$

**Remark.** The proof of Theorem 2 for  $\theta = b$  or  $\phi$  is quite easy, while the proof of it for  $\theta = a$  is relatively difficult.

**Theorem 3.** For all  $\varepsilon > 0$ , if  $F(0) < p_c$ , then

$$\lim_{n \to \infty} -\frac{1}{n} \log P(\theta_{0,n} \le n(\mu - \varepsilon)) = \alpha(\varepsilon, F) > 0.$$

 $\theta = a, b, c \text{ or } \phi.$ 

## **Proof of Theorem 3 for** $\theta = a$

For any  $n \in Z^+$ , let  $X_n := -\log P(a_{0,n} \le n(\mu - \varepsilon))$ .

for all  $m, n \in Z^+$ , note that " $a_{0,m} \leq m(\mu - \varepsilon)$ " and " $a_{m,m+n} \leq n(\mu - \varepsilon)$ " are both **decreasing events**. So using FKG inequality, we get

$$P(a_{0,m} \le m(\mu - \varepsilon))P(a_{m,m+n} \le n(\mu - \varepsilon)) \le P(a_{0,m} \le m(\mu - \varepsilon), a_{m,m+n} \le n(\mu - \varepsilon)).$$

But obviously,  $P(a_{m,m+n} \leq n(\mu - \varepsilon)) = P(a_{0,n} \leq n(\mu - \varepsilon)), \{a_{0,m} \leq m(\mu - \varepsilon), a_{m,m+n} \leq n(\mu - \varepsilon)\} \subseteq \{a_{0,m+n} \leq (m+n)(\mu - \varepsilon)\}.$ 

So we get

$$P(a_{0,m} \le m(\mu - \varepsilon))P(a_{0,n} \le n(\mu - \varepsilon)) \le P(a_{0,m+n} \le (m+n)(\mu - \varepsilon)).$$

It equals to

$$-\log P(a_{0,m+n} \le (m+n)(\mu-\varepsilon)) \le [-\log P(a_{0,m} \le m(\mu-\varepsilon))] + [-\log P(a_{0,n} \le n(\mu-\varepsilon))].$$

Thus  $X_{m+n} \leq X_m + X_n$  holds for any  $m, n \in Z^+$ . Use a small analytical trick we can easily get  $\lim_{n \to \infty} \frac{X_n}{n}$  exists, so

$$\lim_{n \to \infty} -\frac{1}{n} \log P(a0, n \le n(\mu - \varepsilon)) = \alpha(\varepsilon, F) > 0.$$

**Remark.** When  $\theta = b, c$  or  $\phi$ , the proof of Theorem 3 is quite similar.

Next we will check the right tail.

**Theorem 4.(Kesten,1986)** If  $F(0) < p_c$  and  $Ee^{rt(e)} < \infty$  for some r, then for all  $\varepsilon > 0$ , there exist  $c_1$  and  $c_2$ , such that

$$P(\theta_{0,n} \ge n(\mu + \varepsilon)) \le c_1 \exp(-c_2 n).$$

A general case is as follows:

Theorem 5.(Chow & Zhang,2003,Annals of Applied Probability,1601-1614) If  $F(0) < p_c$  and  $Ee^{rt(e)} < \infty$  for some r, then for all  $\varepsilon > 0$ , there exist  $c_1$  and  $c_2$ , such that

$$P(\theta_{0,n} \ge n(\mu + \varepsilon)) \le c_1 \exp(-c_2 n^d).$$

To prove this theorem, we need to prove some lemmas first.

Lemma 3. Define

 $T_{l,k,m} := \inf\{t(\gamma) : \gamma \quad is \quad a \quad path \quad from \quad (l,0,0) \quad to \quad (k,0,0), \quad and \quad \gamma \subseteq [0,k] \times [-m,m]^2\}.$ 

then

$$\lim_{m \longrightarrow \infty} \frac{T_{0,m,m}}{m} = \mu \qquad a.s. \quad and \quad in \quad L_1.$$

**Lemma 4.** If  $F(0) < p_c$ ,  $\varepsilon > 0$  and  $E \exp(rt(e)) < \infty$  for some r > 0. Then there exists a constant  $\eta > 0$ , such that

$$P(T_{0,k,m} \ge k(\mu + \varepsilon)) \le \exp(-\eta k)$$
 for all  $k \ge m$ .

We will take several steps to prove Theorem 5.

## Step 1, Proof of Lemma 3.

Define

 $T_{m,n}(k) := \inf\{t(\gamma): \gamma \quad is \quad a \quad path \quad from \quad (m,0,0) \quad to \quad (n,0,0), \quad and \quad \gamma \subseteq [m-k,n+k] \times Z^{d-1}\}.$ 

By a standard subadditive argument,

$$\lim_{n \to \infty} \frac{t_{0,n}(k)}{n} = \mu(k) = \inf_{n} \frac{Et_{0,n}(k)}{n} \qquad a.s. \quad and \quad in \quad L_1.$$

 $\because \forall n, t_{0,n}(k) \downarrow \text{ as } k \uparrow,$ 

 $\therefore \mu(k) \downarrow \text{ as } k \uparrow.$ 

For fixed k and  $\omega$ , let  $a_{0,n}(\omega) := \lim_{k \to \infty} t_{0,n}(k)(\omega), \ \mu := \lim_{n \to \infty} \frac{a_{0,n}}{n}$ , then obviously,  $a_{0,n} \leq t_{0,n}(k), \ \mu \leq \mu(k)$ .

Fix n,  $\forall \varepsilon > 0$ , take k large such that

$$\mu(k) \le \frac{Et_{0,n}(k)}{n} \le \frac{Ea_{0,n}}{n} + \varepsilon \le \mu + \varepsilon.$$

So for k large, we have  $\mu \leq \mu(k) \leq \mu + \varepsilon$ , which is sufficient to get  $\lim_{n \to \infty} \mu(k) = \mu$ . The next thing for us to do is to compare  $t_{0,k}(0)$  and  $T_{0,k,m}$ .

Let  $\gamma \subseteq [0, n] \times Z^{d-1}$  be a path from (0,0,0) to (n,0,0) with  $T(\gamma) = t_{0,n}(0)$ . Define

$$h_n(\gamma) := \max_{2 \le i \le 3} \{ |m_i| : (m_1, m_2, m_3) \in \gamma \},\$$

and

$$h_n := max\{h_n(\gamma) : \gamma \text{ is a route for } t_{0,n}\}.$$

It is known(see Theorem 8.15 in Smythe and Wierman(1978)) that

$$\limsup_{n \longrightarrow \infty} \frac{h_n}{n} \le 1 \qquad almost \quad surely. \tag{(*)}$$

Let  $H_n := \{\frac{h_n}{n} \le 1\}$ . Then

$$Et_{0,m}(0) \ge E(t_{0,m}(0); H_m) = E(T_{0,m,m}; H_m) = E(T_{0,m,m}) - E(T_{0,m,m}; H_m^c),$$
$$T_{0,m,m} \le \sum_{e \in \gamma} t(e).$$

where  $\gamma$  is the path from (0,0,0) to (m,0,0) along the first coordinate.

By Cauchy-Schwarz inequality, we have:

$$\left[\frac{E(T_{0,m,m};H_m^c)}{m}\right]^2 \le E\left(\frac{\sum\limits_{e \in \gamma} t(e)}{m}\right)^2 \cdot P(H_m^c).$$
  
By (\*) we have  $P(H_m^c) \to 0 (m \to \infty)$ . So  $\lim_{m \to \infty} \frac{E(T_{0,m,m};H_m^c)}{m} = 0.$ 

$$\begin{split} \text{Then } \mu &= \lim_{m \to \infty} \frac{Et_{0,m}(0)}{m} \geq \lim_{m \to \infty} \frac{ET_{0,m,m}}{m}. \\ \text{Note that } \lim_{m \to \infty} \frac{ET_{0,m,m}}{m} \geq \lim_{m \to \infty} \frac{Et_{0,m}(0)}{m} \geq \mu. \quad \text{ So Lemma 3 follows.} \\ \text{Remark. For } M > 0, \text{ let } P(0 \leq t(e) \leq M) = \delta > 0. \end{split}$$

Let A(k) be the event that all those 2k edges from (-k,0,0) to (0,0,0) and from (n,0,0) to (n+k,0,0) along the first coordinate taking values less than M, then

$$\begin{split} P(A(k)) &\geq \delta^{2k} \text{ on } A(k), \qquad t_{-k,n+k}(0) \leq t_{0,n}(k) + 2kM.\\ \text{So } \mu(0) &\leq \frac{Et_{-k,n+k}(0)}{n} = \frac{Et_{0,n+2k}(0)}{n} \leq \frac{Et_{0,n}(k)}{n} + \frac{2kM}{n}.\\ \text{Let } n \to \infty, \text{ we get } \mu(0) \leq \mu(k) \to \mu(k \to \infty). \qquad \therefore \mu(0) \leq \mu.\\ \text{But we have } \mu(k) \downarrow \text{ as } k \uparrow \text{ and } \mu(k) \to \mu(k \to \infty).\\ \text{So we have } \mu(k) = \mu(\forall k), \text{ which is certainly a very strange thing.} \end{split}$$

## Step 2, Proof of Lemma 4 from Lemma 3.

Let k = nm, we can obviously see from the graph that

$$P(T_{0,nm,m} \ge nm(\mu + 2\varepsilon)) \le P(\sum_{i=0}^{n-1} T_{im,(i+1)m,m} \ge nm(\mu + 2\varepsilon)).$$

where  $T_{im,(i+1)m,m}$  are i.i.d., with a common distribution as  $T_{0,m,m}$ . We take m large such that  $ET_{0,m,m} \leq m(\mu + \varepsilon)$ .

By Lemma 3 and a standard large deviation result, we get

$$P(T_{0,nm,m} \ge nm(\mu + 2\varepsilon)) \le P(\sum_{i=0}^{n-1} T_{im,(i+1)m,m} \ge nm(\mu + 2\varepsilon))$$
$$\le P(\sum_{i=0}^{n-1} (T_{im,(i+1)m,m} - ET_{im,(i+1)m,m}) \ge nm\varepsilon) \le \exp(-Cnm\varepsilon),$$

where C > 0 is a constant. So Lemma 4 follows.

Step 3, Proof of Theorem 5 from Lemma 4 for  $\theta = \phi$ .

Take k = mn and divide  $[0, k]^2$  into  $(\frac{k}{m})^2 = n^2$  equal subsquares of size  $m \times m$ , which are called  $S_1, S_2, \dots, S_{n^2}$ .

Since  $\{\phi([0,k] \times S_i) \ge k(\mu + \varepsilon)\}$  and  $\{\phi([0,k] \times S_j) \ge k(\mu + \varepsilon)\}$  are independent for  $i \ne j$ , and  $\{\phi_{0,k} \ge k(\mu + \varepsilon)\} \subseteq \bigcap_{i=1}^{n^2} \{\phi([0,k] \times S_i) \ge k(\mu + \varepsilon)\}$ , we then have  $B(\ne i \ge k(\mu + \varepsilon)) \le [B(\phi([0,k] \times [0,m]^2) \ge k(\mu + \varepsilon))]^{n^2}$  (1)

$$P(\phi_{0,k} \ge k(\mu + \varepsilon)) \le [P(\phi([0,k] \times [0,m]^2) \ge k(\mu + \varepsilon))]^{n^2}.$$
 (1)  
By Lemma 4 and translation invariance, we have

 $P(\phi([0,k] \times [0,m]^2) \ge k(\mu + \varepsilon)) \le P(T_{0,k,\frac{m}{2}} \ge k(\mu + \varepsilon)) \le \exp(-\eta k).$ (2) Combining (1) and (2), we get that for k large and m|k,

$$P(\phi_{0,k} \ge k(\mu + \varepsilon)) \le \exp(-\eta k n^2) = \exp(-C'k^3),$$

where  $C' = \frac{\eta}{m^2}$ . So Theorem 5 follows.

A further result is as follows.

Theorem 5.(Chow & Zhang,2003,Annals of Applied Probability,1601-1614) If  $F(0) < p_c$  and  $Ee^{rt(e)} < \infty$  for some r, then for all  $\varepsilon > 0$ , there exists a constant  $\beta(\varepsilon, F, d)$ , such that

$$\lim_{n \to \infty} \frac{-1}{n^d} \log P(\phi_{0,n} \ge n(\mu + \varepsilon)) = \beta(\varepsilon, F, d).$$